

Handout 1: Spectrum of a Gaussian Random Matrix

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Our goal in this handout is to give upper and lower bounds on the singular values of a Gaussian random matrix. Before we proceed, let us recall some basic definitions and results concerning the singular values of a matrix.

Let $\mathbf{A} \in \mathbb{R}^{m \times n}$ be a given matrix. The i -th singular value of $\mathbf{A}$ can be computed as $s_i(\mathbf{A}) = \sqrt{\lambda_i(\mathbf{A}^T \mathbf{A})}$, where $\lambda_i(\mathbf{A}^T \mathbf{A})$ is the i -th eigenvalue of $\mathbf{A}^T \mathbf{A}$. The spectral norm (i.e., the largest singular value) of $\mathbf{A}$ is given by

$$\|\mathbf{A}\| = \max_{\mathbf{v} \in \mathbb{S}^{n-1}} \|\mathbf{A}\mathbf{v}\|_2 = \max_{\mathbf{u} \in \mathbb{S}^{m-1}, \mathbf{v} \in \mathbb{S}^{n-1}} \mathbf{u}^T \mathbf{A} \mathbf{v}, \quad (1)$$

where $\mathbb{S}^{n-1} = \{\mathbf{v} \in \mathbb{R}^n : \|\mathbf{v}\|_2 = 1\}$. Observe that when $m = n$ and $\mathbf{A}$ is symmetric, we have

$$\|\mathbf{A}\| = \max_{\mathbf{u} \in \mathbb{S}^{n-1}} |\mathbf{u}^T \mathbf{A} \mathbf{u}|.$$

This can be established by considering the spectral decomposition of $\mathbf{A}$.

Suppose that $m \geq n$. It is easy to show that for any $t \in (0, 1)$, if $\|\mathbf{A}^T \mathbf{A} - \mathbf{I}_n\| \leq t$, then $1 - t \leq s_i(\mathbf{A}) \leq 1 + t$ for $i \in \{1, \dots, n\}$. This gives us a handle to bound all the singular values of $\mathbf{A}$.

Now, let $\mathbf{X} \in \mathbb{R}^{m \times n}$ be a matrix whose entries are independent and identically distributed (iid) standard Gaussian random variables; i.e., $X_{ij} \sim \mathcal{N}(0, 1)$ for $i \in \{1, \dots, m\}$ and $j \in \{1, \dots, n\}$. Observe that $\mathbb{E}[\mathbf{X}^T \mathbf{X}] = m \mathbf{I}_n$. Hence, our goal of bounding the singular values of $\mathbf{X}$ can be achieved by bounding

$$\left\| \frac{1}{m} \mathbf{X}^T \mathbf{X} - \mathbf{I}_n \right\|.$$

1 Computing Spectral Norm on a Net

Recall from (1) that the spectral norm of a matrix involves the maximization of a certain quantity over the sphere $\mathbb{S}^{n-1}$. In the study of random matrices, it is often useful to perform the maximization over a *finite* subset $\mathcal{N}$ of $\mathbb{S}^{n-1}$. To get a small approximation error, the set $\mathcal{N}$ should cover $\mathbb{S}^{n-1}$ sufficiently well. This motivates the following definition:

Definition 1 (ϵ -Net) We say that $\mathcal{N}$ is an ϵ -net of $\mathbb{S}^{n-1}$ if (i) $\mathcal{N} \subseteq \mathbb{S}^{n-1}$ and (ii) for every $\mathbf{v} \in \mathbb{S}^{n-1}$, there exists a $\mathbf{v}_0 \in \mathcal{N}$ such that $\|\mathbf{v} - \mathbf{v}_0\|_2 \leq \epsilon$.

By considering suitable nets on $\mathbb{S}^{m-1}$ and $\mathbb{S}^{n-1}$, we have the following result:

Proposition 1 (Bilinear Form on a net) Let $\mathbf{A} \in \mathbb{R}^{m \times n}$ be a given matrix. Let $\mathcal{M}$ be a δ -net of $\mathbb{S}^{m-1}$ and $\mathcal{N}$ be an ϵ -net of $\mathbb{S}^{n-1}$, where $\delta, \epsilon \in [0, 1)$. Then,

$$\|\mathbf{A}\| \leq \frac{1}{(1 - \delta)(1 - \epsilon)} \max_{\mathbf{u} \in \mathcal{M}, \mathbf{v} \in \mathcal{N}} \mathbf{u}^T \mathbf{A} \mathbf{v}.$$

Proof Let $\bar{\mathbf{v}} \in \mathbb{S}^{n-1}$ be such that $\|\mathbf{A}\| = \|\mathbf{A}\bar{\mathbf{v}}\|_2$. By definition of a net, we can find a $\bar{\mathbf{v}}_0 \in \mathcal{N}$ such that $\|\bar{\mathbf{v}} - \bar{\mathbf{v}}_0\|_2 \leq \epsilon$. Hence, we have

$$\|\mathbf{A}\| = \|\mathbf{A}(\bar{\mathbf{v}}_0 + (\bar{\mathbf{v}} - \bar{\mathbf{v}}_0))\|_2 \leq \|\mathbf{A}\bar{\mathbf{v}}_0\|_2 + \|\mathbf{A}\| \cdot \|\bar{\mathbf{v}} - \bar{\mathbf{v}}_0\|_2 = \|\mathbf{A}\bar{\mathbf{v}}_0\|_2 + \epsilon\|\mathbf{A}\|,$$

which implies that

$$\|\mathbf{A}\| \leq \frac{1}{1-\epsilon} \|\mathbf{A}\bar{\mathbf{v}}_0\|_2. \quad (2)$$

Now, let $\bar{\mathbf{u}} \in \mathbb{S}^{m-1}$ be such that $\|\mathbf{A}\bar{\mathbf{v}}_0\|_2 = \bar{\mathbf{u}}^T \mathbf{A}\bar{\mathbf{v}}_0$. Again, by definition of a net, we can find a $\bar{\mathbf{u}}_0 \in \mathcal{M}$ such that $\|\bar{\mathbf{u}} - \bar{\mathbf{u}}_0\|_2 \leq \delta$. Hence, we have

$$\|\mathbf{A}\bar{\mathbf{v}}_0\|_2 = \bar{\mathbf{u}}_0^T \mathbf{A}\bar{\mathbf{v}}_0 + (\bar{\mathbf{u}} - \bar{\mathbf{u}}_0)^T \mathbf{A}\bar{\mathbf{v}}_0 \leq \max_{\mathbf{u} \in \mathcal{M}, \mathbf{v} \in \mathcal{N}} \mathbf{u}^T \mathbf{A}\mathbf{v} + \delta\|\mathbf{A}\bar{\mathbf{v}}_0\|_2. \quad (3)$$

Putting (2) and (3) together yields the desired result. $\square$

In the case where $\mathbf{A} \in \mathbb{R}^{n \times n}$ is symmetric, we have the following result:

Proposition 2 (Quadratic Form on a net) *Let $\mathbf{A} \in \mathbb{R}^{n \times n}$ be a given symmetric matrix. Let $\mathcal{N}$ be an ϵ -net of $\mathbb{S}^{n-1}$, where $\epsilon \in [0, 1/2)$. Then,*

$$\|\mathbf{A}\| \leq \frac{1}{1-2\epsilon} \max_{\mathbf{u} \in \mathcal{N}} |\mathbf{u}^T \mathbf{A}\mathbf{u}|.$$

Proof Let $\bar{\mathbf{u}} \in \mathbb{S}^{n-1}$ be such that $\|\mathbf{A}\| = |\bar{\mathbf{u}}^T \mathbf{A}\bar{\mathbf{u}}|$ and $\bar{\mathbf{u}}_0 \in \mathcal{N}$ be such that $\|\bar{\mathbf{u}} - \bar{\mathbf{u}}_0\|_2 \leq \epsilon$. Noting that $\|\mathbf{A}\| = \max_{\mathbf{u} \in \mathbb{S}^{n-1}} \|\mathbf{A}\mathbf{u}\|_2$ and $\|\bar{\mathbf{u}}_0\|_2 = 1$ by definition, we compute

$$\begin{aligned} \|\mathbf{A}\| &= |\bar{\mathbf{u}}^T \mathbf{A}\bar{\mathbf{u}}| \leq |\bar{\mathbf{u}}^T \mathbf{A}(\bar{\mathbf{u}} - \bar{\mathbf{u}}_0)| + |\bar{\mathbf{u}}^T \mathbf{A}\bar{\mathbf{u}}_0| \\ &\leq \epsilon\|\mathbf{A}\bar{\mathbf{u}}\|_2 + |(\bar{\mathbf{u}} - \bar{\mathbf{u}}_0)^T \mathbf{A}\bar{\mathbf{u}}_0| + |\bar{\mathbf{u}}_0^T \mathbf{A}\bar{\mathbf{u}}_0| \\ &\leq \epsilon\|\mathbf{A}\| + \|\mathbf{A}(\bar{\mathbf{u}} - \bar{\mathbf{u}}_0)\|_2 + \max_{\mathbf{u} \in \mathcal{N}} |\mathbf{u}^T \mathbf{A}\mathbf{u}| \\ &\leq 2\epsilon\|\mathbf{A}\| + \max_{\mathbf{u} \in \mathcal{N}} |\mathbf{u}^T \mathbf{A}\mathbf{u}|. \end{aligned}$$

This completes the proof. $\square$

Next, let us bound the size of an ϵ -net of $\mathbb{S}^{n-1}$ using a volume argument.

Proposition 3 (Size of a Net) *Let $\epsilon \in (0, 1)$ be given. Then, there exists an ϵ -net $\mathcal{N}$ of $\mathbb{S}^{n-1}$ with*

$$|\mathcal{N}| \leq \left(\frac{2}{\epsilon} + 1\right)^n - \left(\frac{2}{\epsilon} - 1\right)^n.$$

Proof Let $\mathcal{N}$ be a maximal cardinality subset of $\mathbb{S}^{n-1}$ such that for any distinct $\mathbf{u}, \mathbf{v} \in \mathcal{N}$, we have $\|\mathbf{u} - \mathbf{v}\|_2 > \epsilon$. By the maximality of $\mathcal{N}$, we see that $\mathcal{N}$ is an ϵ -net of $\mathbb{S}^{n-1}$. Now, observe that $B(\mathbf{u}, \epsilon/2) \cap B(\mathbf{v}, \epsilon/2) = \emptyset$ for every distinct $\mathbf{u}, \mathbf{v} \in \mathcal{N}$, where $B(\mathbf{x}, \epsilon)$ denotes the Euclidean ball centered at $\mathbf{x} \in \mathbb{R}^n$ with radius ϵ . Moreover, it is clear that

$$\bigcup_{\mathbf{u} \in \mathcal{N}} B(\mathbf{u}, \epsilon/2) \subseteq B(\mathbf{0}, 1 + \epsilon/2) \setminus B(\mathbf{0}, 1 - \epsilon/2).$$

Hence, by comparing the volumes of the balls, we have

$$|\mathcal{N}| \cdot \text{vol}(B(\mathbf{0}, \epsilon/2)) \leq \text{vol}(B(\mathbf{0}, 1 + \epsilon/2)) - \text{vol}(B(\mathbf{0}, 1 - \epsilon/2)). \quad (4)$$

Now, recall that

$$\text{vol}(B(\mathbf{x}, \epsilon)) = \epsilon^n \text{vol}(B(\mathbf{0}, 1)) \quad \text{for any } \mathbf{x} \in \mathbb{R}^n.$$

Hence, upon dividing both sides of (4) by $\text{vol}(B(\mathbf{0}, 1))$, we obtain

$$|\mathcal{N}| \cdot \left(\frac{\epsilon}{2}\right)^n \leq \left(1 + \frac{\epsilon}{2}\right)^n - \left(1 - \frac{\epsilon}{2}\right)^n.$$

This implies the desired result. $\square$

2 Concentration Inequality for χ^2 Random Variable

Let $\mathcal{N}$ be an ϵ -net of $\mathbb{S}^{n-1}$. Using Proposition 2, we have

$$\left\| \frac{1}{m} \mathbf{X}^T \mathbf{X} - \mathbf{I}_n \right\| \leq \frac{1}{1 - 2\epsilon} \max_{\mathbf{u} \in \mathcal{N}} \left| \mathbf{u}^T \left(\frac{1}{m} \mathbf{X}^T \mathbf{X} - \mathbf{I}_n \right) \mathbf{u} \right| = \frac{1}{1 - 2\epsilon} \max_{\mathbf{u} \in \mathcal{N}} \left| \frac{1}{m} \|\mathbf{X}\mathbf{u}\|_2^2 - 1 \right|. \quad (5)$$

Now, using the fact that X_{ij} are iid according to $\mathcal{N}(0, 1)$ for $i \in \{1, \dots, m\}$, $j \in \{1, \dots, n\}$ and $\|\mathbf{u}\|_2 = 1$ for any $\mathbf{u} \in \mathcal{N}$, we have

$$\frac{1}{m} \|\mathbf{X}\mathbf{u}\|_2^2 - 1 = \frac{1}{m} \sum_{i=1}^m (g_i^2 - 1),$$

where $g_1, \dots, g_m$ are iid according to $\mathcal{N}(0, 1)$. As an aside, the random variable $\sum_{i=1}^m g_i^2$ follows the so-called χ^2 -distribution with m degrees of freedom. The above development motivates us to establish the following result:

Proposition 4 (Concentration Inequality for χ^2 Random Variable) *Let $g_1, \dots, g_m$ be iid according to $\mathcal{N}(0, 1)$. Then, for any $t \in [0, 1]$,*

$$\Pr \left(\left| \frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \right| \geq t \right) \leq 2 \exp(-mt^2/8).$$

Proof Note that

$$\Pr \left(\left| \frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \right| \geq t \right) \leq \Pr \left(\frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \geq t \right) + \Pr \left(\frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \leq -t \right). \quad (6)$$

Hence, it suffices to bound the two terms on the right-hand side of (6) separately. To bound the first term in (6), observe that by Markov's inequality, for any $\lambda \geq 0$, we have

$$\begin{aligned} \Pr \left(\frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \geq t \right) &= \Pr \left[\exp \left(\lambda \sum_{i=1}^m (g_i^2 - 1) \right) \geq \exp(\lambda mt) \right] \\ &\leq \exp(-\lambda mt) \cdot \mathbb{E} \left[\exp \left(\lambda \sum_{i=1}^m (g_i^2 - 1) \right) \right]. \end{aligned} \quad (7)$$

Since $g_1, \dots, g_m$ are iid, we have

$$\mathbb{E} \left[\exp \left(\lambda \sum_{i=1}^m (g_i^2 - 1) \right) \right] = \prod_{i=1}^m \mathbb{E} [\exp(\lambda(g_i^2 - 1))] = (\mathbb{E} [\exp(\lambda(g_1^2 - 1))])^m.$$

Using the density function of a standard Gaussian random variable, for any $\lambda < 1/2$, we have

$$\begin{aligned}\mathbb{E} [\exp (\lambda(g_1^2 - 1))] &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp(\lambda(x^2 - 1)) \cdot \exp(-x^2/2) dx \\ &= \sigma \cdot \exp(-\lambda) \cdot \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} \exp(-x^2/2\sigma^2) dx,\end{aligned}$$

where $\sigma = (1 - 2\lambda)^{-1/2}$. Upon noting that

$$x \mapsto \frac{1}{\sqrt{2\pi}\sigma} \exp(-x^2/2\sigma^2)$$

is the density function of a Gaussian random variable with mean 0 and variance σ^2 and using the inequality $-x - x^2 \leq \ln(1 - x)$, which is valid for any $x \in [0, 1/2]$, we obtain, for any $\lambda \in [0, 1/4]$,

$$\mathbb{E} [\exp (\lambda(g_1^2 - 1))] = \frac{\exp(-\lambda)}{\sqrt{1 - 2\lambda}} \leq \exp(2\lambda^2).$$

Substituting the above into (7), we get, for any $\lambda \in [0, 1/4]$,

$$\Pr \left(\frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \geq t \right) \leq \exp(m(2\lambda^2 - \lambda t)).$$

Since the above bound holds for any $\lambda \in [0, 1/4]$, we can get the best bound by minimizing the right-hand side over λ . In particular, by setting $\lambda = t/4$ and noting that $\lambda \in [0, 1/4]$ whenever $t \in [0, 1]$, we obtain

$$\Pr \left(\frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \geq t \right) \leq \exp(-mt^2/8).$$

To bound the second term in (6), we proceed in a similar manner. Specifically, by Markov's inequality, for any $\lambda \geq 0$, we have

$$\Pr \left(\frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \leq -t \right) \leq \exp(-\lambda mt) \cdot (\mathbb{E} [\exp (\lambda(1 - g_1^2))])^m.$$

Since the inequality $(1 + 2x)^{-1/2} \leq \exp(2x^2 - x)$ holds for all $x \geq 0$, a simple calculation yields

$$\mathbb{E} [\exp (\lambda(1 - g_1^2))] = \frac{\exp(\lambda)}{\sqrt{1 + 2\lambda}} \leq \exp(2\lambda^2).$$

Using the same argument as before, we obtain, for any $t \in [0, 1]$,

$$\Pr \left(\frac{1}{m} \sum_{i=1}^m (g_i^2 - 1) \leq -t \right) \leq \exp(-mt^2/8).$$

This completes the proof. □

3 Bounding the Spectrum of Gaussian Random Matrix

With the results in the previous sections, we are now ready to give upper and lower bounds on the singular values of a Gaussian random matrix. Let $\mathcal{N}$ be a $(1/4)$ -net of $\mathbb{S}^{n-1}$. By (5) and Propositions 3 and 4, we have

$$\begin{aligned} \Pr\left(\left\|\frac{1}{m}\mathbf{X}^T\mathbf{X} - \mathbf{I}_n\right\| \geq 2t\right) &\leq \Pr\left(\max_{\mathbf{u} \in \mathcal{N}} \left|\frac{1}{m}\|\mathbf{X}\mathbf{u}\|_2^2 - 1\right| \geq t\right) \leq \sum_{\mathbf{u} \in \mathcal{N}} \Pr\left(\left|\frac{1}{m}\|\mathbf{X}\mathbf{u}\|_2^2 - 1\right| \geq t\right) \\ &\leq 2 \cdot 9^n \cdot \exp(-mt^2/8). \end{aligned}$$

By setting $t^2 = \frac{8 \ln 9}{m}(n + \eta^2)$ for some $\eta \geq 0$ and noting that $\sqrt{a+b} \leq \sqrt{a} + \sqrt{b}$ for any $a, b \geq 0$, we get

$$\Pr\left(\left\|\frac{1}{m}\mathbf{X}^T\mathbf{X} - \mathbf{I}_n\right\| \geq 2 \cdot \sqrt{\frac{8 \ln 9}{m}} \cdot (\sqrt{n} + \eta)\right) \leq 2 \exp(-(\ln 9)\eta^2).$$

It follows that when $\eta > 0$ is sufficiently large, we have

$$\sqrt{m} - c(\sqrt{n} + \eta) \leq s_i(\mathbf{X}) \leq \sqrt{m} + c(\sqrt{n} + \eta) \quad (8)$$

for $i \in \{1, \dots, n\}$ with high probability, where $c = 2\sqrt{8 \ln 9}$.

4 Remarks

1. It is possible to establish upper and lower bounds on the singular values of more general classes of random matrices; see [3, Chapter 4] for a development in this direction.
2. Note that the lower bound in (8) becomes less useful as m tends to n ; i.e., the matrix $\mathbf{X}$ becomes more square. To estimate the least singular value of an almost-square Gaussian random matrix, one needs more sophisticated machinery. We refer the interested reader to the paper [1] and survey [2] for further reading.

References

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